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# **SOME PSYCHOMETRIC CHARACTERISTICS OF THE FAMILY ADAPTATION SCALE (FAS) ON A SAMPLE OF YOUNG PEOPLE IN SERBIA**

The Family Adaptation Scale has already been translated into Serbian and used on a sample of families in our country. Despite its frequent usage in Serbia, its factorial structure and predictive value has never been explored. According to theoretical background, FAS was composed of items from two aspects of adaptation (inner and external) and global appraisal, but they are loading one latent variable. In this research, we investigated factorial structure and reliability of the scale, as well as its predictive validity in relation to some aspects

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of risky behaviour. We performed these analyses in two samples: high-school and university students. Using the Principal component method, we extracted only one factor in both samples. Reliability analysis showed very high internal consistency. However, predictive validity is very small and suspicious.

**KEYWORDS:** Family adaptation scale, psychometric characteristics, high-school and university students.

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## INTRODUCTION

During the last two decades, we can notice the trend of examination of mental health and good functioning through the indicators of the so-called positive mental health. In this sense, the most commonly used term is subjective well-being (Stanojević, 2012). It refers to “the individual's cognitive or affective assessment of one's own life“ (Diener & al., 2002, p. 63). Cognitive assessment is primarily concerned with the concept of satisfaction (satisfaction with life, satisfaction with marriage, job satisfaction, etc.) while the emotional aspect implies the frequency with which people experience pleasant (such as joy) or unpleasant emotions (such as depression) (Diener & al., 1997). From this point, we can conclude that there are three basic components of subjective well-being (Diener & al., 1997): satisfaction, pleasant emotions, and low levels of unpleasant emotions. These three components constitute one global factor of interconnected variables. Each of these components can be separated into parts. Global satisfaction can be divided into satisfaction with different domains of an individual's life: recreation, love, marriage, friendship. Pleasant emotions can be divided into specific feelings: joy, love, pride. Unpleasant feelings can be separated into specific emotions: shame, guilt, sorrow, anger, anxiety. Each of these parts can be divided into even smaller parts (Diener & al., 1997).

Subjective well-being is usually measured by self-reported techniques. Initially, satisfaction measurements of life or hap-

piness were made up of only one item by which the respondent should determine the overall assessment of one's own life (Diener & al., 2002). More recently, scales with a larger number of items have appeared, the most famous of which is the Satisfaction with Life Scale (Diener & al., 1985). This scale has been translated into many of languages and used in many surveys. The construct of "life satisfaction" was examined in relation to many other constructs, in particular with indicators of disease and health (Stanojević, 2019). In various studies, a negative correlation was obtained between satisfaction with life and depression (Blais & al., 1989; Guney & al., 2010; Koivumaa-Honkanen & al., 2004; Saunders & Roy, 1999, Stanojević & al., 2014), anxiety and general psychological distress (Arrindell & Ettema, 1986, according to Pavot & Diener, 1993), anxiety and helplessness (Guney & al., 2010) and a negative affect (Larsen & al., 1985; Smead, 1991, according to Pavot & Diener, 1993), while positive correlation was obtained with a positive affect (George, 1991), self-esteem, and quality of family relationships (Raboteg-Šarić & al., 2009). Satisfaction with life was also examined in relation to the personality traits (Hosseinkhanzadeh & Taher, 2013), affective temperament (Jaredić & al., 2017), as well as many other variables.

The presence of a positive and negative affect, as an emotional component of subjective well-being, was assessed by various instruments, among which, perhaps, the most popular are the scales PANAS (Watson & al., 1988) and PANAS-X (Watson & Clark, 1994). As might be expected, research has shown that negative affection was positively, and positive affection negatively correlated with measures of anxiety, depression, distress and dysfunction (Crawford & Henry, 2004; Watson & al., 1988).

## SATISFACTION WITH FAMILY

In an effort to determine factors associated with quality family life, researchers have focused on variables related to the satisfaction with family. Satisfaction with family has been directly related to many other well-being variables, both global and family-specific, like family cohesion, communication, adaptability, and other indicators of family life quality (Poff, Zabriskie, Townsend, 2010).

However, tradition of family satisfaction researches is not too long. Zabriskie and Ward (2013) claimed it began in the 1970s with the Family Life Questionnaire (Guerney, 1977), while the 1980s brought a variety of new approaches to the measurement of family satisfaction. The most known was Olson's (1979) Circumflex Model of Marital and Family Systems. The scale that was based on that approach was Family Adaptability and Cohesion Evaluation Scale (FACES), with several revisions in past few decades. In the 1980s McCollum and associates developed well-known scale named Kansas Family Life Satisfaction Questionnaire (McCollum, Schumm, & Russell, 1988; Schumm, McCollum, Bugaighis, Jurich, & Bollman, 1986). This scale measures family satisfaction based on a differential approach and asks how satisfied or dissatisfied the respondents are with specific relationships (i.e., marital, parental, relationship among siblings). Also, respondents are asked to assess global satisfaction with family relationship. Another scale was developed with aim to provide measure of global satisfaction, but with respondent's family of origin. It is a Likert-type Family Satisfaction Scale (Carver & Jones, 1992) with 20 items. With aim to encompass affective aspect of satisfaction with life and family as a domain, Barraca, Yarto, and Olea (2000) developed the Family Satisfaction by Adjectives Scale. That scale consisted of 27 items and was widely used for measuring affective component of family satisfaction.

## FAMILY ADAPTATION SCALE (FAS)

Family Adaptation Scale was developed at the end of 1980s (Antonovsky & Sourani, 1988). Using this scale, each family member estimates how satisfied they are with the family's adaptability in the environment and how satisfied they are with the family, independently from objective indicators that can indicate that certain maladaptive behaviours are present. The scale consists of 10 items to which the respondents answer by estimating on a scale of 1 (Not satisfied at all) to 7 (Completely satisfied). In original study (Antonovsky & Sourani, 1988) authors did not offer the answer "completely dissatisfied", but we choose to use same answers as in many studies in the Serbian sample. The theoretical range of scores is from 10 to 70, so the higher score achieved on the scale indicates higher family satisfaction, and the lower indicates less satisfaction with the family. The five items of this scale are related to satisfaction with the inner adaptation of the family, two to the family's adaptability to the environment in which they live, and the remaining three are general and capture both of these aspects of adaptation. We did not use additional item for global appraisal of satisfaction with family life, due to our research aim and simplicity of analyses we aimed for. The reliability of the scale measured by Cronbach's  $\alpha$ , on the sample of adult respondents from Israel, was 0,87 (Antonovsky & Sourani, 1988).

The Family Adaptation Scale has already been translated into Serbian and used on a sample of families in our country. Despite its frequent usage (Jaređić, 2016; Mitić, 1997; Minić & al., 2011; Pavićević & Krstić, 2013; Pavićević & Minić, 2011a; Pavićević & Minić, 2011b; Stanojević, 2019) its factorial structure and predictive value has never been explored. According to theoretical background, FAS was composed of items from two aspects of adaptation (inner and external) and global ap-

praisal, but they are loading one latent variable (Antonovsky & Sourani, 1988).

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## METHOD

Main aim of this study was to investigate some psychometric characteristics of Family Adaptation Scale (Antonovsky & Sourani, 1988) in two samples: high-school students and university students. Namely, we wanted to explore factorial structure and reliability of the scale, as well as its predictive validity in relation to some aspects of risky behaviour of young people. To explore factorial structure of FAS, we performed Explorative Factor Analyses (EFA); reliability was expressed with Cronbach's  $\alpha$ ; predictive validity was tested by calculating significant differences between respondents who reported different types of risky behaviour.

## MEASURES

Main measure in this research is Family Adaptation Scale (Antonovsky & Sourani, 1988). This scale is already described in detail.

We measured risky behaviour with several questions about: smoking cigarettes (never, sometimes, on regular bases), drinking alcohol (never, sometimes, on regular bases), taking drugs (never, sometimes, on regular bases), practicing unsafe sex (never, sometimes, on regular bases), getting into fights (never, sometimes, on regular bases) and stealing (yes/no).

## SAMPLE

The sample consisted of 816 respondents, divided in two groups: high-school (N=434 or 53,2%) and university students (N=382 or 46,8%). Samples were uniform in percentage of males/females, 51% of females in high-school and 52% in university student sample. Average age of high-school student

was 17,3 and of university students 20,6. Respondents were informed about the research and anonymity of collected data was guaranteed. Data were collected in schools and faculties in Central Serbia and Kosovo and Metohija.

## RESULTS

Firstly, we performed two separated explorative factor analyses, Principal Component Method, so we could get an overview in structure of FAS (Antonovsky & Sourani, 1988) in both samples. As a minimal value of factor loadings we used ,03. To assess suitability of data for factor analyses, we performed Kaiser-Meyer-Olkin Measure of Sampling Adequacy and Bartlett's Test of Sphericity. Results of these tests are shown in Table 1.

TABLE 1: KAISER-MEYER-OLKIN AND BARTLETT'S TEST

KAISER-MEYER-OLKIN MEASURE OF SAMPLING ADEQUACY.		HIGH-SCHOOL	UNIVERSITY
		,933	,939
BARTLETT'S TEST OF SPHERICITY	APPROX. CHI-SQUARE	3279,84	3530,33
	DF	45	45
	SIG.	,000	,000

Kaiser-Meyer-Olkin and Bartlett's Test showed that our data are suitable for factor analyses procedure. Principal Component Analysis showed that, in both samples, we extracted only one component with eigenvalue above 1,00. That factor can explain about 65% (in high school student sample) or 71% (in university student sample) of variance of satisfaction with family. These data are shown in Table 2. and Table 3, correspondingly.

TABLE 2: TOTAL VARIANCE EXPLAINED (HIGH SCHOOL STUDENT SAMPLE)

COMPONENT	INITIAL EIGENVALUES			EXTRACTION SUMS OF SQUARED LOADINGS		
	TOTAL	% OF VARIANCE	CUMULATIVE %	TOTAL	% OF VARIANCE	CUMULATIVE %
1	6,480	64,803	64,803	6,480	64,803	64,803
2	,739	7,391	72,194			
3	,662	6,625	78,819			
4	,495	4,948	83,767			
5	,383	3,828	87,595			
6	,371	3,712	91,307			
7	,263	2,625	93,932			
8	,231	2,314	96,246			
9	,194	1,945	98,191			
10	,181	1,809	100,000			

Extraction Method: Principal Component Analysis.

TABLE 3: TOTAL VARIANCE EXPLAINED (UNIVERSITY STUDENT SAMPLE)

COMPONENT	INITIAL EIGENVALUES			EXTRACTION SUMS OF SQUARED LOADINGS		
	TOTAL	% OF VARIANCE	CUMULATIVE %	TOTAL	% OF VARIANCE	CUMULATIVE %
1	7,145	71,453	71,453	7,145	71,453	71,453
2	,637	6,369	77,822			
3	,524	5,238	83,060			
4	,374	3,744	86,804			
5	,287	2,868	89,672			
6	,258	2,578	92,250			
7	,244	2,439	94,688			
8	,210	2,096	96,785			

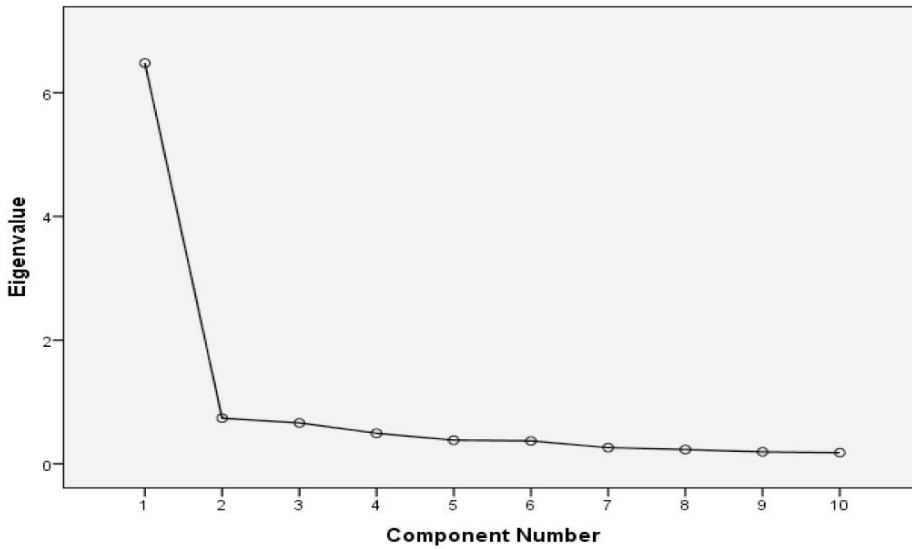


TABLE 3: TOTAL VARIANCE EXPLAINED (UNIVERSITY STUDENT SAMPLE)

9	,184	1,838	98,623			
10	,138	1,377	100,000			
<b>EXTRACTION METHOD: PRINCIPAL COMPONENT ANALYSIS.</b>						

Scree plot also indicated that FAS item are loading at one factor (Graphics 1. and 2).

**Scree Plot**



GRAPHIC 1: SCREE PLOT FOR HIGH-SCHOOL SAMPLE

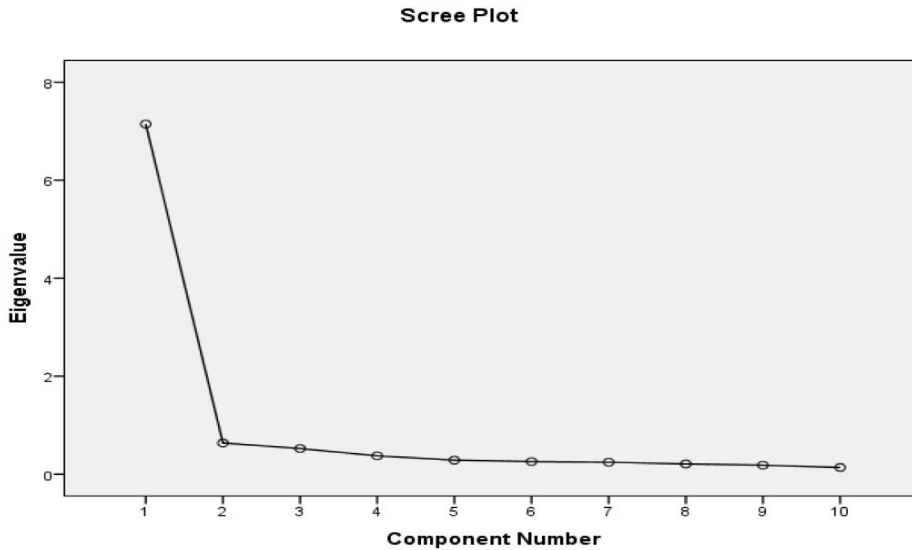
In Table 4. we presented data about item loadings. We can see that they are pretty high, and that they are somewhat higher in university students sample.

TABLE 4: COMPONENT MATRIX FOR BOTH SAMPLES

	COMPONENT 1	
	HIGH-SCHOOL STUDENTS SAMPLE	UNIVERSITY STUDENTS SAMPLE
ITEM 3	,864	,885
ITEM 10	,855	,860
ITEM 5	,845	,855
ITEM 9	,843	,854
ITEM 7	,833	,853
ITEM 2	,829	,850
ITEM 1	,808	,850
ITEM 4	,755	,841
ITEM 6	,738	,827
ITEM 8	,656	,773
Extraction Method: Principal Component Analysis.		

## DESCRIPTIVE ANALYSIS

With aim to provide complete descriptive data about FAS, we firstly performed Kolmogorov-Smirnov test, its statistics was significant for both samples (Table 5), meaning that our data significantly step out of normal distribution.



GRAPHIC 2: SCREE PLOT FOR UNIVERSITY SAMPLE

TABLE 5: KOLMOGOROV-SMIRNOV TEST SEPARATELY FOR HIGH-SCHOOL AND UNIVERSITY STUDENTS

	KOLMOGOROV-SMIRNOV		
	STATISTIC	DF	SIG.
HIGH-SCHOOL STUDENTS	,199	434	,000
UNIVERSITY STUDENTS	,212	382	,000

However, Kolmogorov-Smirnov test is usually significant whenever we have large samples, which is the case in the present research. Therefore, we were looking at skewness and kurtosis (Table 6) whose ranges are above normal distributions values.

TABLE 6: DESCRIPTIVE STATISTICS

SAMPLE	N	MIN	MAX	M	SD	SKEWNESS	KURTOSIS
HIGH-SCHOOL STUDENTS	434	10,00	70,00	59,94	11,93	-1,991	4,100
HIGH-SCHOOL STUDENTS	382	10,00	70,00	59,80	12,75	-2,025	3,919

## RELIABILITY ANALYSIS

We tested reliability by calculated Cronbach`s  $\alpha$ . In our research, obtained coefficients were  $\alpha=,94$  (high-school sample) and  $\alpha=,95$  (university sample). Total-item correlations were moderate to very high (high-school sample) and very high (university sample). Those data are shown in Table 7.

TABLE 7: ADDITIONAL DATA IN RELIABILITY ANALYSES

	ITEM-TOTAL CORRELATION		CRONBACH'S ALPHA IF ITEM DELETED	
	(HIGH-SCHOOL SAMPLE)	(UNIVERSITY SAMPLE)	(HIGH-SCHOOL SAMPLE)	(UNIVERSITY SAMPLE)
ITEM 3	,749	,809	,930	,949
ITEM 10	,772	,808	,929	,949
ITEM 5	,817	,803	,926	,948
ITEM 9	,698	,780	,932	,949
ITEM 7	,800	,819	,927	,948
ITEM 2	,684	,730	,934	,952
ITEM 1	,792	,828	,928	,947
ITEM 4	,590	,809	,937	,948

TABLE 7: ADDITIONAL DATA IN RELIABILITY ANALYSES

ITEM 6	,794	,850	,928	,947
ITEM 8	,812	,812	,927	,948

## PREDICTIVE VALIDITY

Predictive validity of FAS was tested by calculating significant differences between respondents who reported different types of risky behaviour. We measured: smoking cigarettes, drinking alcohol, taking drugs (light and heavy separately), practicing unsafe sex, getting into fights and stealing.

As we can see in Table 8 we have not obtained any differences between respondents who do not smoke cigarettes, smoke sometimes, and smoke on regular bases, in both samples.

TABLE 8: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DON`T SMOKE CIGARETTES, SMOKE SOMETIMES, AND SMOKE ON REGULAR BASES (KRUSKAL WALLIS TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	295	221,98	CHI-SQ.	1,20	NEVER	241	193,03	CHI-SQ.	,47
SOMETIMES	56	209,75	DF	2	SOMETIMES	59	195,24	DF	2
REGULARLY	83	206,81	As. SIG.	,55	REGULARLY	82	184,30	As. SIG.	,79

Table 9. shows that we have not obtained any differences in satisfaction with family between respondents who do not smoke cigarettes, smoke sometimes, and smoke on regular bases, in both samples.

TABLE 9: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DO NOT DRINK ALCOHOL, DRINK SOMETIMES, AND DRINK ON REGULAR BASES (KRUSKAL WALLIS TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	80	236,16	CHI-SQ.	5,10	NEVER	76	205,38	CHI-SQ.	1,534
SOMETIMES	323	216,80	DF	2	SOMETIMES	285	188,31	DF	2
REGULARLY	31	176,61	As. SIG.	,08	REGULARLY	21	184,52	As. SIG.	,46

Although we offered three answers on question about taking light drugs, our respondents choose one of two answers: never and sometimes. So, we tested potential differences between those two groups. Table 10. shows that significant differences were not obtained, in both samples.

TABLE 10: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DO NOT TAKE LIGHT DRUGS, TAKE THEM SOMETIMES, AND TAKE THEM ON REGULAR BASES (MANN-WHITNEY U TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	364	193,04	M-W U	4940,5	NEVER	364	193,04	M-W U	2351,5
SOMETIMES	17	147,32	WILCOX. W	5405,5	SOMETIMES	17	147,32	WILCOX. W	2504,5
			Z	-1,56				Z	-1,68
			ASYMP. SIG.	,119				ASYMP. SIG.	,09

We obtained same type of answers when we asked about heavy drugs, so our respondent grouped in two groups, again. But this time, we obtained significant differences in high-school student sample (Table 11). Respondents who did not take heavy drugs had higher satisfaction with family. In

university student sample, significant differences were not found.

TABLE 11: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DO NOT TAKE HEAVY DRUGS, TAKE THEM SOMETIMES, AND TAKE THEM ON REGULAR BASES (MANN-WHITNEY U TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	432	218,41	M-W U	38,50	NEVER	364	193,04	M-W U	2351,50
SOMETIMES	2	20,75	WILCOX. W	41,50	SOMETIMES	17	147,32	WILCOX. W	2504,50
			Z	-2,229				Z	-1,68
			ASYMP. SIG.	,03				ASYMP. SIG.	,09

When we talk about getting into fights, Kruskal Wallis test showed that there were not significant differences in family satisfaction between respondents in both samples (Table 12).

TABLE 12: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DO NOT FIGHT, DO IT SOMETIMES, AND DO IT ON REGULAR BASES (KRUSKAL WALLIS TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	324	221,40	CHI-SQ.	3,355	NEVER	326	194,54	CHI-SQ.	1,744
SOMETIMES	101	211,18	DF	2	SOMETIMES	52	174,70	DF	2
REGULARLY	9	147,94	As. SIG.	,187	REGULARLY	4	162,25	As. SIG.	,42

However, when we look at satisfaction with family in groups of respondents who steal and do not steal, we found significant differences in samples, high-school and university students (Table 13). Respondents who did not steal had significant higher satisfaction with family.

TABLE 13: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO STEAL AND DO NOT STEAL (MANN-WHITNEY U TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
YES	82	181,99	M-W U	11520,50	YES	11	147,32	M-W U	1155,000
NO	328	211,38	WILCOX. W	14923,50	NO	367	193,04	WILCOX. W	1221,000
			Z	-2,01				Z	-2,426
			ASYMP.	,044				ASYMP. SIG.	,015

Kruskal-Wallis test showed that there were significant differences in family satisfaction between groups of respondents who do not practice unsafe sex, do it sometimes, and do it on regular bases, but only in university students sample (Table 14).

TABLE 14: SIGNIFICANT DIFFERENCES BETWEEN RESPONDENTS WHO DON`T PRACTICE UNSAFE SEX, DO IT SOMETIMES AND DO IT ON REGULAR BASE (KRUSKAL WALLIS TEST)

HIGH-SCHOOL SAMPLE					UNIVERSITY SAMPLE				
	N	MEAN RANK				N	MEAN RANK		
NEVER	315	220,43	CHI-SQ.	,695	NEVER	199	214,42	CHI-SQ.	18,114
SOMETIMES	91	208,13	DF	2	SOMETIMES	135	165,21	DF	2
REGULARLY	28	215,02	As. SIG.	,707	REGULARLY	48	170,40	As. SIG.	,000

Then, we performed three Man-Witney tests to determine between which groups, in the university sample, we obtained those differences. These tests showed that respondents who do not practice unsafe sex have significantly higher satisfaction with family then those who do it sometimes ( $Z = -4,03$ ,



$p < ,01$ ), and do it on regular bases ( $Z = -2,46$ ,  $p < ,01$ ). There was no significant difference between the last two groups.

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## DISCUSSION

The main aim of this study was to investigate some psychometric characteristics of Family Adaptation Scale (Antonovsky & Sourani, 1988) in two samples: high-school students and university students.

The results we obtained indicated that we can use Family Adaptation Scale in both samples. Factorial analyses showed that, in both samples, all items of FAS are loading at one factor, with high percent of variance. This is an important result because the scale was developed with five items which are related to satisfaction with the inner adaptation of the family, two to the family's adaptability to the environment in which they live, and the remaining three are general and capture both of these aspects of adaptation (Antonovsky & Sourani, 1988). So, theoretically, we could expect one, two or three factors. However, our results showed, without any doubts, that we can talk about one latent variable.

Reliability analysis indicated that FAS is a highly reliable instrument for high-school ( $\alpha = ,94$ ) and university students ( $\alpha = ,95$ ). In earlier studies, reliability of FAS was also very high, but somewhere smaller than in our research. In parental sample Cronbach's  $\alpha$  was ,874 (Antonovsky & Sourani, 1988), while in student sample it was ,864 (Pavicevic & Krstić, 2013).

To test normal distribution of FAS, we performed a few tests: Kolmogorov-Smirnov, skewness, and kurtosis. Kolmogorov-Smirnov test was significant in both samples, and skewness and kurtosis values were significantly above normal distribution values. Kurtosis values are even out of ranges that are allowed to be used for confirmative analysis and structural equation modelling analysis (Tabachnick & Fidell, 2001). In

the original study, Antonovsky and Sourani (1988) noticed respondents' proneness to choose extreme positive answers, and they did not even offer the answer "completely dissatisfied". Although they did not report new distribution data, it was an attempt to make kurtosis somewhat lower. However, it is usual in Likert type of scales to have same levels of positive and negative answers, and the Serbian version of FAS was used in that way. However, we cannot recommend data obtained with this scale to be analysed in parametric tests, confirmative analysis, and structural equation modelling analysis. However, several non-parametric tests can be used in future research.

According to the suggestions above, we performed non-parametric analyses to test significant differences in family satisfaction between groups of respondents who are prone to some forms of risky behaviour. We expected to obtain different FAS scores in groups with different level of risky behaviour. That would be a certain evaluation of FAS predictive validity, at least related to risky behaviour of young people. However, we did not obtain significant differences in satisfaction with family between groups of respondents (in both samples) who reported that they do not smoke cigarettes, drink alcohol, take light or heavy drugs, and gets into fights, and those who reported that they are prone to this types of risky behaviour sometimes or regularly. On the other side, we obtained significant differences, in both samples, when we compared groups of respondents who steal and do not steal. Respondents who reported that they steal had lower satisfaction with life, in both samples. Also, in university sample, we obtained significant difference between groups of respondents who practice unsafe sex. Respondents who never practice unsafe sex had significantly higher satisfaction with family then those who do it sometimes and do it on regular base. In summary, we can conclude that we cannot be sure in predictive validity of FAS, especially for risky behaviour. We may count on its predictive

value for certain types of risky behaviour in university student sample. For high-school students, its predictive value is very small or non-existent. We can explain this later result if we remember that high-school students are in middle of adolescence, which is characterized as life period when young people are turning more from family to peers. Also, this is life phase when the young want to try many things and are more prone to risky behaviour. However, we have to conclude that predictive validity of FAS is pretty small, especially in high-school sample.

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## CONCLUSION

Based on our data, we can make several highlights: FAS is reliable scale with clear one-factor structure; FAS significantly steps out of normal distribution and its data can be analysed only with non-parametric tests. Predictive validity is very small and somewhat better in university than in high-school sample.

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**ДРАГАНА З. СТАНОЈЕВИЋ**

УНИВЕРЗИТЕТ У ПРИШТИНИ СА ПРИВРЕМЕНИМ СЕДИШТЕМ  
У КОСОВСКОЈ МИТРОВИЦИ, ФИЛОЗОФСКИ ФАКУЛТЕТ

**ОЛИВЕРА Б. РАДОВИЋ**

УНИВЕРЗИТЕТ У ПРИШТИНИ СА ПРИВРЕМЕНИМ СЕДИШТЕМ У КОСОВСКОЈ МИТРОВИЦИ,  
ФИЛОЗОФСКИ ФАКУЛТЕТ

**МИРОСЛАВ Ж. КРСТИЋ**

УНИВЕРЗИТЕТ У ПРИШТИНИ СА ПРИВРЕМЕНИМ СЕДИШТЕМ  
У КОСОВСКОЈ МИТРОВИЦИ, ФИЛОЗОФСКИ ФАКУЛТЕТ

РЕЗИМЕ

### НЕКЕ ПСИХОМЕТРИЈСКЕ КАРАКТЕРИСТИКЕ СКАЛЕ ПОРОДИЧНЕ АДАПТАЦИЈЕ НА УЗОРКУ МЛАДИХ У СРБИЈИ

Главни циљ ове студије био је испитати неке психометријске Скале породичне адаптације (Antonovsky & Sourani, 1988) на узорцима средњошколаца и студената. Наиме, желели смо да испитамо факторску структуру и поузданост ове скале, као и њену предиктивну валидност у односу на неке аспекте ризичног понашања младих. Да бисмо утврдили факторску структуру скале, спровели смо две одвојене експлоративне факторске анализе (ЕФА); поузданост је изражена Кронбаховим  $\alpha$  коефицијентом, а предиктивна валидност испитана је израчунавањем значајности разлика између група које су пријавиле различите нивое ризичног понашања.

Анализа главних компоненти показала је, у оба узорка, да се може екстраховати само један фактор карактеристичне вредности изнад 1,00 и да се њиме може објаснити око 65% (код средњошколаца) и 71% (код студената) варијансе задовољства породицом. Поузданост је у оба узорка била врло висока  $\alpha = ,94$  (код средњошколаца) и  $\alpha = ,95$  (код студената). Тотал-ајтем корелације кретале су се од умерених до врло високих (код средњошколаца), односно у опсегу високих и врло високих (код студената). Тестови нормалности дистрибуције били су значајни у оба узорка и показали су да подаци значајно одступају од нормалне дистрибуције. Вредности куртозиса чак прелазе и најшире дефинисане вредности за спровођење конфирмативне факторске